

Predicting stand structure using limited measurements

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Outline of the presentation

Introduction

Methods

- Linear Prediction

- Examples of BLP

Predicting the distribution of diameters

- Percentile-based diameter distribution

- Distributions of order statistics

- BLP in this case

- The effect of order statistics on prediction accuracy

Discussion

The inventory of a single stand

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- ▶ In this study, these are described by the total basal area, diameter distribution and height-diameter curve, each of which is expressed by tree species.
- ▶ The main aim of this study was to develop tools for producing a stand description from ground-measured data
- ▶ Because of limited measurement resources for a single stand, the available information is very limited and the use of measurement information should be as effective as possible.

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Group2: Sample measurements (of response)

- ▶ Sample tree height(s)
- ▶ Sample order statistic(s) of the angle count sample plot(s), called quantile trees.

*The use is based on **linear prediction**.*

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In applications, we usually use estimated $\boldsymbol{\mu}$'s and \mathbf{V} 's to obtain an Estimated BLP (EBLP).

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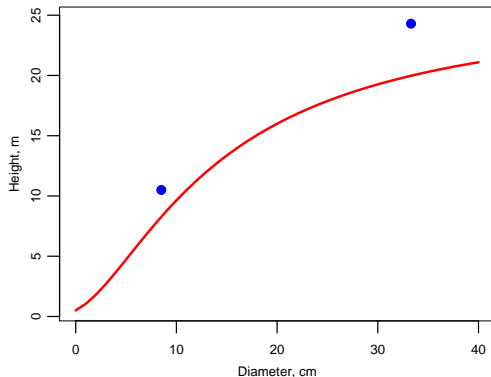
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- ▶ In spatial statistics, kriging is a method where unobserved responses, \mathbf{x}_1 at spatial locations are predicted using observed responses, \mathbf{x}_2 at spatial locations. Matrices \mathbf{V} are obtained from an estimated variogram.

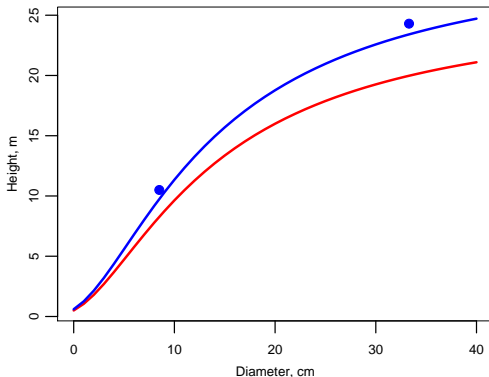
Predicting Height-Diameter curve

The fixed part of a mixed model gives the expected height for a tree with a given diameter. If we have sample tree heights measured, we can predict the random effects of the model (x_1) using sample tree heights (x_2)



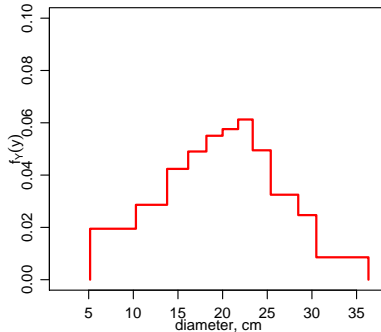
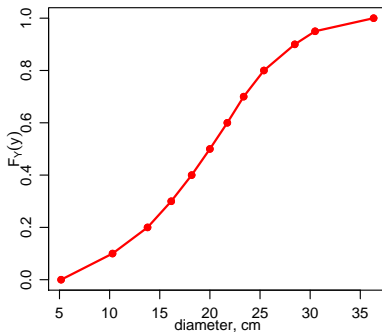
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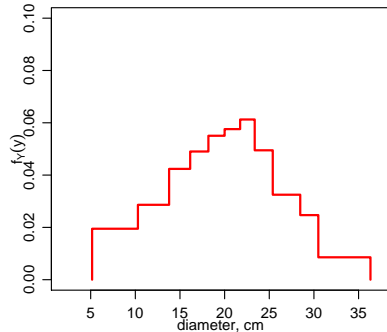
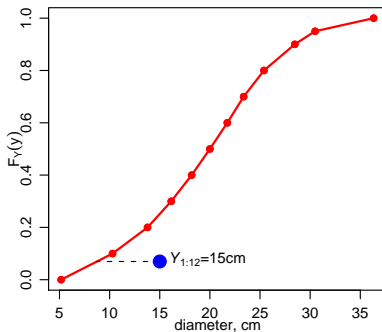
Predicting the diameter distribution

- ▶ I predict diameter percentiles using models that are estimated from some data *a priori*.



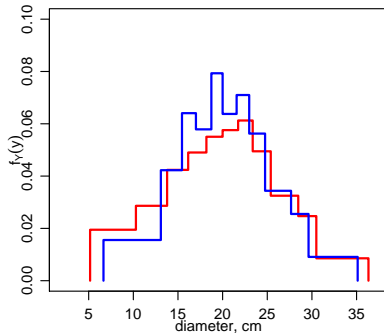
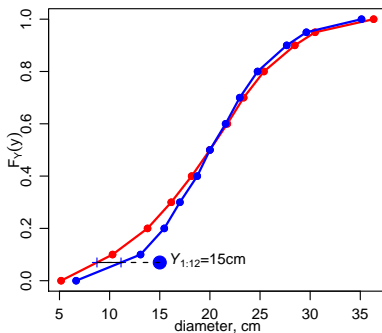
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- ▶ The quantile tree is interpreted as a measured percentile,
- ▶ so we can use it (x_2) to predict the residuals of our percentile models (x_1).



Percentile-based diameter distribution

Certain percentiles in stand k are modeled using Group 1 variables

$$\begin{aligned}
 d_{0\%,k} &= E(d_{0\%}|DGM_k, G_k, T_k, \dots) + e_{0\%,k} \\
 d_{10\%,k} &= E(d_{10\%}|DGM_k, G_k, T_k, \dots) + e_{10\%,k} \\
 &\vdots \\
 d_{90\%,k} &= E(d_{90\%}|DGM_k, G_k, T_k, \dots) + e_{90\%,k} \\
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The same in matrix form:

$$\mathbf{d}_k = E(\mathbf{d}|DGM_k, G_k, T_k, \dots) + \mathbf{e}_k,$$

where residuals \mathbf{e}_k are actually stand effects with nondiagonal $\text{var}(\mathbf{e}_k) = \mathbf{D}$. We will drop index k hereafter.

Sample order statistics

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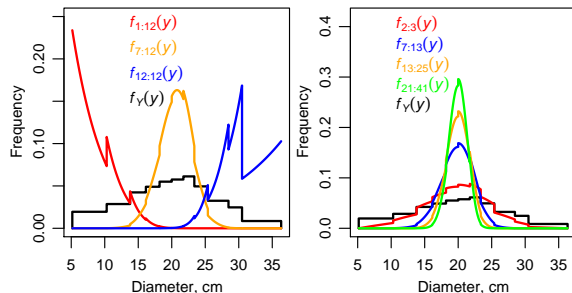
$$f_{r:n}(y) = \frac{n!}{(r-1)!(n-r)!} f_Y(y) [F_Y(y)]^{r-1} [1 - F_Y(y)]^{n-r}$$

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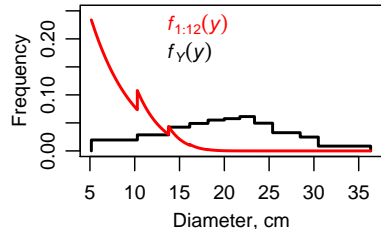


Examples of distributions, black line is the underlying population distribution

Expectation and variance of $Y_{r:n}$

Recall the example, where $Y_{1:12}$ was observed

Assuming that the predicted $f_Y(y)$ is the diameter distribution, we know the **distribution of $Y_{1:12}$** .



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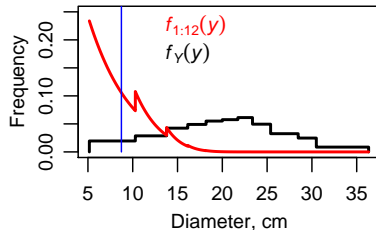
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$$E(Y_{1:12}) = 8.73\text{cm and}$$

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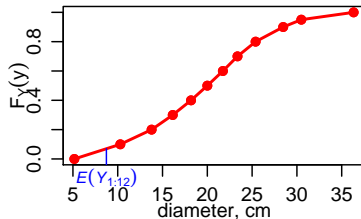
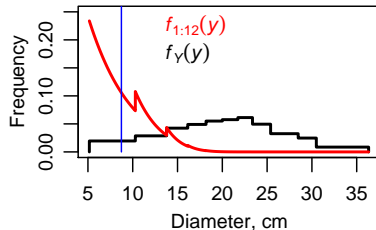
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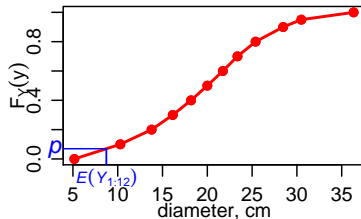
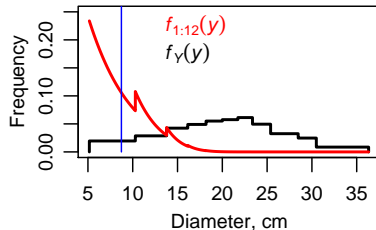
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The model for the measured percentile

Recall the model for the predefined percentiles

$$\mathbf{d}_k = E(\mathbf{d}|DGM, G, T, \dots) + \mathbf{e}_k,$$

where $\text{var}(\mathbf{e}_k) = \mathbf{D}$.

The model for this observation is

$$d_{7\%} + \epsilon = E(d_{7\%}|DGM, G, T, \dots) + e_{7\%} + \epsilon,$$

where ϵ is the measurement error of of the percentile, with $\text{var}(\epsilon) = \text{var}(Y_{1:12}) = 8.09$.

We approximate $E(d_{7\%}|DGM, G, T, \dots)$, $\text{var}(e_{7\%})$ and $\text{cov}(\epsilon, e_{7\%})$ by interpolating linearly $E(\mathbf{d}|DGM, G, T, \dots)$ and \mathbf{D} for the 7th percentile.

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We can now write

$$\begin{bmatrix} \mathbf{d} \\ d_{7\%} + \epsilon \end{bmatrix} \sim \left(\begin{bmatrix} E(\mathbf{d}|DGM, G, T, \dots) \\ E(d_{7\%}|DGM, G, T, \dots) \end{bmatrix}, \begin{bmatrix} \mathbf{D} & \text{cov}(\mathbf{e}, e_{7\%}) \\ \text{cov}(e_{7\%}, \mathbf{e}) & \text{var}(e_{7\%}) + \text{var}(\epsilon) \end{bmatrix} \right),$$

where all except for \mathbf{d} is known.

BLP of stand effects

$$\begin{bmatrix} \mathbf{d} \\ d_{7\%} + \epsilon \end{bmatrix} \sim \left(\begin{bmatrix} E(\mathbf{d}|DGM, G, T, \dots) \\ E(d_{7\%}|DGM, G, T, \dots) \end{bmatrix}, \begin{bmatrix} \mathbf{D} & \text{cov}(\mathbf{e}, \mathbf{e}_{7\%}) \\ \text{cov}(\mathbf{e}_{7\%}, \mathbf{e}) & \text{var}(\mathbf{e}_{7\%}) + \text{var}(\epsilon) \end{bmatrix} \right),$$

The BLP is calculated as

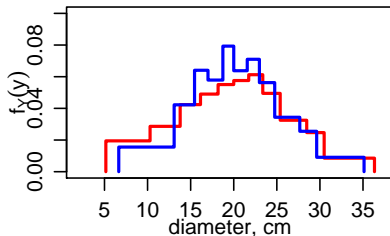
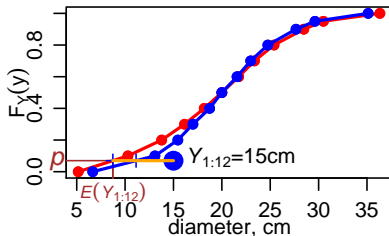
$$\hat{\mathbf{d}} = E(\mathbf{d}|DGM, G, T, \dots) + \text{cov}(\mathbf{e}, \mathbf{e}_{7\%})(\text{var}(\mathbf{e}_{7\%}) + \text{var}(\epsilon))^{-1}(d_{7\%} - E(d_{7\%}|DGM, G, T, \dots))$$

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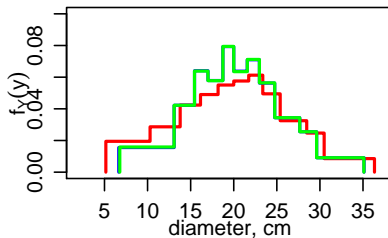
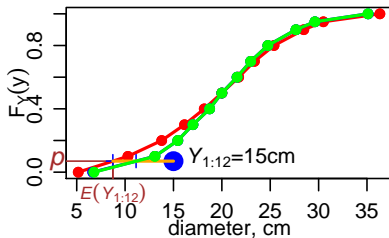
Oops

- ▶ Our $f_{r:n}(y)$ was based on assumption that $E(\mathbf{d}|DGM, G, T, \dots)$ gives the *true* distribution of the stand.
- ▶ Now we have a better estimate for the true distribution, namely $\hat{\mathbf{d}}$.
- ▶ That is why we use $\hat{\mathbf{d}}$ to calculate p and $\text{var}(Y_{r:n})$ again, and predict new $\hat{\mathbf{d}}$.

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After 7 iterations the final values were $\rho = 0.064$, $E(Y_{1:12}) = 10.82\text{cm}$ and $\text{var}(Y_{r:n}) = 8.28$.



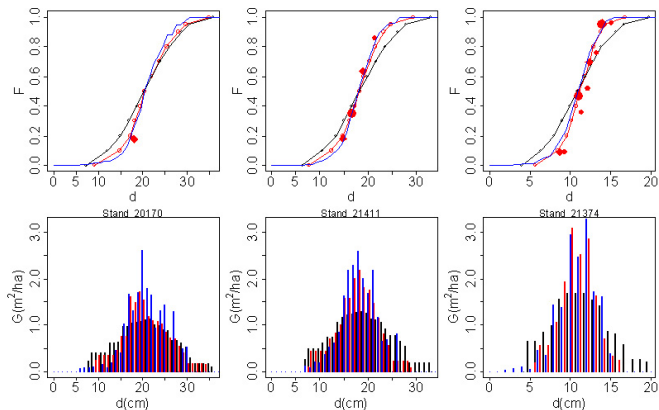
Several observations per stand

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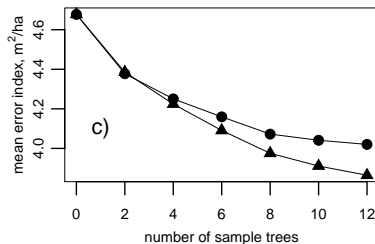
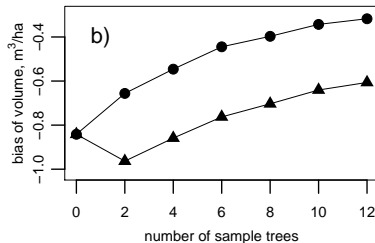
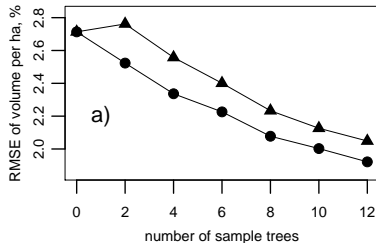
- ▶ If all measurements are from different plots, their “measurement errors” (ϵ 's) are uncorrelated.
- ▶ In the case of several observation per plot, ϵ 's are correlated. Joint distribution of two order statistics is needed to calculate the correlation:

Examples of predictions (True trees from true plots)



True distribution, prediction based on *DGM*, *G*, *T*, *soil* and the prediction based on *DGM*, *G*, *T*, *soil* and sample order statistics (the marks). Same mark is used for measurements from same plot.

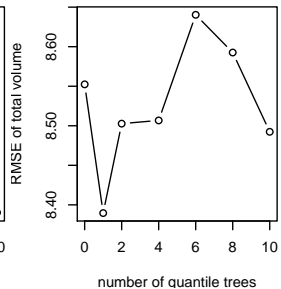
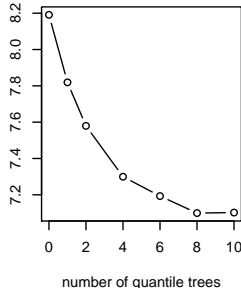
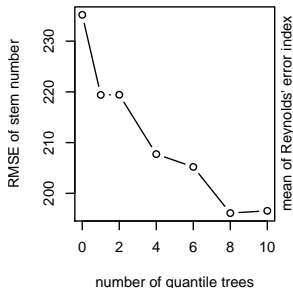
True trees from simulated plots



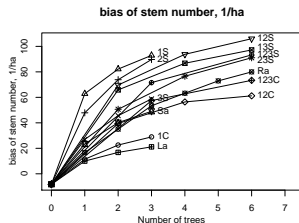
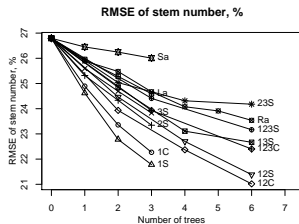
- ▶ Order statistics are taken from *random samples* of Norway spruce stands.
- ▶ Triangles show the prediction after first step, the circle after converged iteration.

True trees from true plots

- ▶ Trees were randomly selected from 3 angle count sample plots
- ▶ Stand variables (DGM, G, N) included measurement errors
- ▶ Observations in a sample seem not to be independent observations from the population



Which quantile to measure?



Quantile trees were selected according to different strategies from 1-3 angle count plots

- ▶ 1S, 2S, 3S: 1st, 2nd and 3rd smallest tree of each plot
- ▶ 12S, 13S, 23S: 1st and 2nd, 1st and 3rd, and 2nd and 3rd smallest trees of each plot
- ▶ 1C, 2C, 3C: 1st, 2nd and 3rd closest tree of each plot
- ▶ 12C, 13C, 23C: 1st and 2nd, 1st and 3rd, and 2nd and 3rd closest trees of each plot
- ▶ Sa: Tree closest to sawtimber limit (17 cm in diameter)
- ▶ La: Largest tree of each plots
- ▶ Ra: Randomly selected trees

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- ▶ Using sample information combined with models provides possibilities to control the accuracy of a forest inventory according to the information needs
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- ▶ Using extremes seems to cause a bias problem

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Publications

Mehtätalo, L. 2005. Localizing a predicted diameter distribution using sample information. *Forest Science* 51(4): 292-302.

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